



EDUCATION

Challenges for Value-Added Assessment of Teacher Effects

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Value-Added Models of Student Achievement

- **Goal of value-added models (VAM) is to estimate causal effects of individual schools and teachers on student growth**
 - **Assumes schools and teachers contribute to growth**
 - **Recognizes that level of achievement is highly dependent on non-educational inputs**
 - **Attempts to separate schooling from other inputs to achievement (e.g., family or student background)**
- **VAM uses statistical analysis of repeated scores from students to estimate teacher effects**

Policy Makers are Very Interested in VAM

- Several states provide school districts with VAM results, including TN, PA, OH, IA
 - Except for TN, use of VAM is within the last couple of year****
- More states are considering implementation (LA, DE, NY and CA)**
- Advocates have proposed using VAM results for teacher promotion, retention, and salary decisions**

High Expectations for VAM

- ❑ “[VAM] provides unbiased estimates of the effects of schooling on individual and group academic progress” (PSDE, 2002)
- ❑ “[VAM] separates annual growth [in scores] into two parts: that which can be attributed to the teacher and that which can be attributed to the student” (Hershberg, 2004)
- ❑ “[VAM] may yield answers to questions such as: Do some teachers demonstrate consistently greater/less effectiveness in adding to student assessment gains?” (PSDE, 2002)

Goals for Today's Talk

- Investigate four challenges to estimating teacher effects using VAM
 - Persistence of teacher effects
 - Effects of student background variables
 - Missing scores and missing teacher links
 - Modeling approach
- Using variety of methods
 - Empirical analyses of data from a large urban school district
 - Analytic analyses
 - Simulation study

Outline

- ❑ **Persistence of teacher effects**
- ❑ Effects of student background variables
- ❑ Missing scores and missing teacher links
- ❑ Modeling approach

A Model for Longitudinal Test Score Data (The Persistence Model)

$$y_{i1} = \mu_1 + \beta_1' \mathbf{x}_{i1} + \lambda_{i1}' \boldsymbol{\eta}_1 + \phi_{i1}' \boldsymbol{\theta}_1 + \epsilon_{i1}$$

$$y_{i2} = \mu_2 + \beta_2' \mathbf{x}_{i2} + \lambda_{i2}' \boldsymbol{\eta}_2 + \phi_{i2}' \boldsymbol{\theta}_2 + \\ \omega_{21} \lambda_{i1}' \boldsymbol{\eta}_1 + \alpha_{21} \phi_{i1}' \boldsymbol{\theta}_1 + \epsilon_{i2}$$

$$y_{i3} = \mu_3 + \beta_3' \mathbf{x}_{i3} + \lambda_{i3}' \boldsymbol{\eta}_3 + \phi_{i3}' \boldsymbol{\theta}_3 + \\ \omega_{32} \lambda_{i2}' \boldsymbol{\eta}_2 + \alpha_{32} \phi_{i2}' \boldsymbol{\theta}_2 + \omega_{31} \lambda_{i1}' \boldsymbol{\eta}_1 + \alpha_{31} \phi_{i1}' \boldsymbol{\theta}_1 + \epsilon_{i3}$$

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Other Models Are Special Cases of This Model

□ The layered model (TVAAS)

- No school effects

- No covariates

- $\alpha_{tt'} \equiv 1$, for all $t > t'$

□ Cross-classified models

- Residual error terms specified by student random growth curves (e.g., random intercepts and slopes)

- $\alpha_{tt'} \equiv 1$, for all $t > t'$

How Do Teacher Effects Persist Over Time?

- Common models assume teacher effects persist undiminished into future years of testing
 - Layered model, cross-classified models
- Our previous explorations found persistence parameters to be significantly less than 1
 - α ranged from 0.2 to 0.3 after one year and 0.1 after two years
 - Very small sample of four schools and 678 students

Fitting the Persistence Model Is Computationally Challenging

- **Covariance matrix of the vector of scores is large and sparse**
 - **Covariance matrix is not block-diagonal**
 - **Cannot be inverted using simple formulas**

- **Hierarchical model packages will fit some cross-classified models but not with persistence**
 - **Challenging to specify even the layered model in MLWin**
 - **Impossible in SAS**
 - **Recent updates to R and HLM will fit cross-classified models but not with persistence parameters**

Bayesian Approach Simplifies Computation

- Does not require inverting the complete covariance matrix
- Markov Chain Monte Carlo (MCMC) algorithm draws from posterior by taking sequential samples from conditional distributions
 - Conditional distributions become simple regression-like problems

Implementing the Bayesian Approach

□ WinBUGS

■ Freeware available at

<http://www.mrc-bsu.cam.ac.uk/bugs/welcome.shtml>

■ Sufficiently fast for moderate problems

■ Too slow for large data sets

□ Developed our own MCMC algorithm and implemented it in C

■ Acceptably fast for even large data sets

Application to a Large Urban School District

- ❑ One of the nation's largest urban school districts
 - Enrollment of about 75,000 students
 - Estimating teacher effects for about 1,500 teachers
- ❑ Spring testing for 1998 to 2002
- ❑ Math and reading scores
- ❑ Students in grades 1 to 5 for the tested years
- ❑ Links students across years, to teachers and to schools
- ❑ Focus on grade 1 teachers in 1998 to grade 5 teachers in 2002
- ❑ 9,295 students in the analysis sample

Persistence Parameters Are Significantly Less Than One

	Math		Reading	
	Mean	Std. Dev.	Mean	Std. Dev.
α_{21}	0.11	0.02	0.15	0.02
α_{31}	0.13	0.02	0.12	0.02
α_{32}	0.14	0.02	0.20	0.03
α_{41}	0.08	0.02	0.11	0.02
α_{42}	0.07	0.02	0.13	0.03
α_{43}	0.09	0.02	0.14	0.03
α_{51}	0.08	0.02	0.10	0.02
α_{52}	0.09	0.02	0.16	0.03
α_{53}	0.08	0.02	0.13	0.03
α_{54}	0.20	0.03	0.23	0.03

Layered and Persistence Models Yield Different Inferences for Some Teachers

- ❑ Correlation between the estimated teacher effects ranges across grades from .73 to .87 for math and .77 to .84 for reading
- ❑ Teacher variance components are much larger for the persistence model (1.3 to 6.8 times larger)
- ❑ Persistence model yields more teachers whose 90% posterior intervals exclude zero, especially for grades 2 and 3
 - Example: grade 3 math,
 - ❑ For persistence model, 105 of the 306 teachers significantly above average teacher
 - ❑ For the layered model only 57 classified as above average
 - ❑ 50 overlap

Outline

- Persistence of teacher effects
- **Effects of student background variables**
- Missing scores and missing teacher links
- Modeling approach

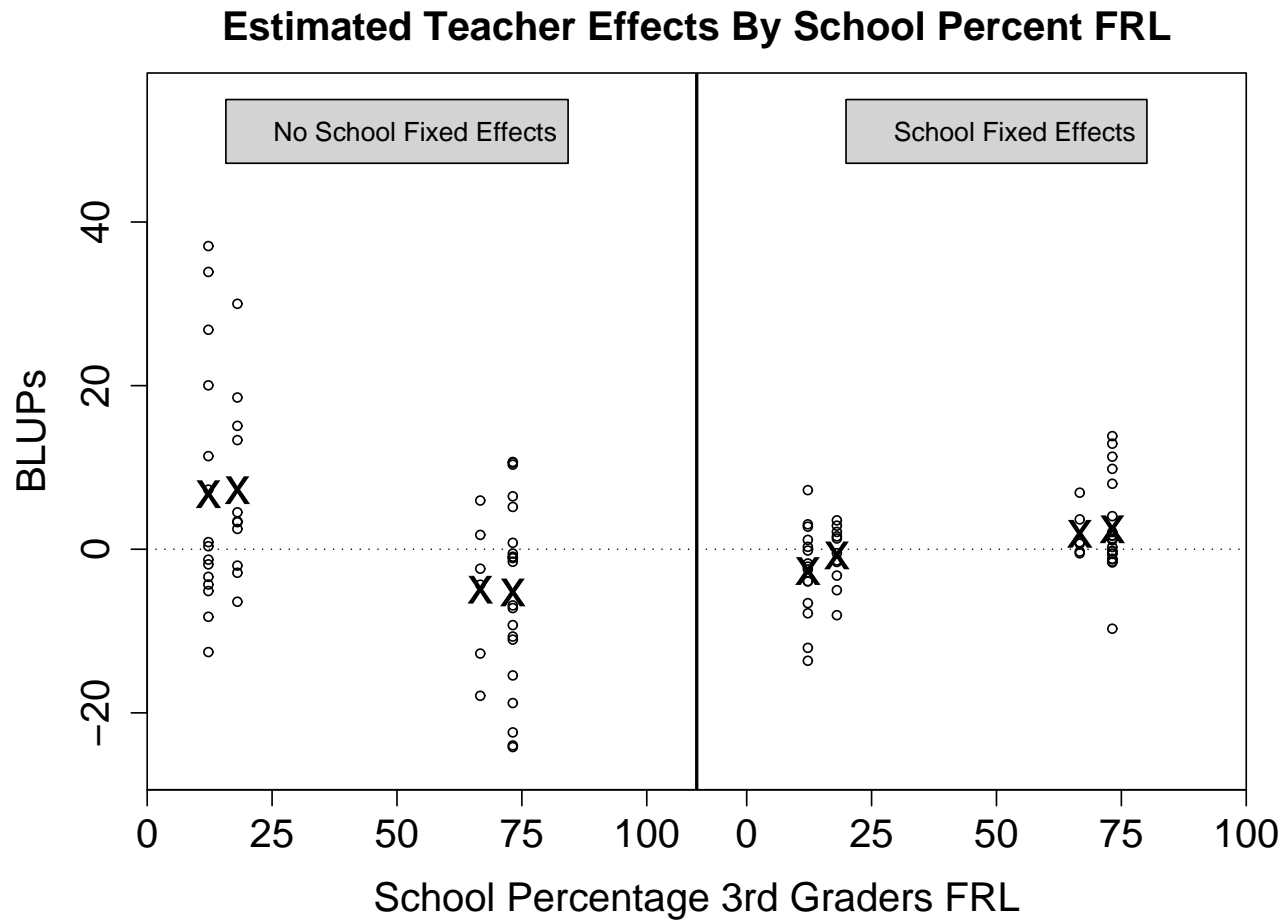
Do Estimated Effects Depend on the Students?

- ❑ Purported value of VAM is it separates teacher effects from student effects**
- ❑ On average, are some estimated teacher effects higher or lower because of the students they teach?**
- ❑ Do estimated effects correlate with student characteristics?**

Estimated Effect Will Depend on Students But Intra-Student Correlation Mitigates Some Errors

- ❑ Students are not uniformly grouped**
- ❑ Some of the inter-classroom heterogeneity is associated with estimated teacher effects**
- ❑ Accounting for intra-student correlation in the multiple scores from a student removes much of inter-classroom heterogeneity if students mix across years**
- ❑ Accounting for intra-student correlation does *not* mitigate confounding when student populations are stratified according to achievement**

Example Where Estimated Teacher Effects Depend on the School



School Level Effects Remain an Issue in Urban School Data

- ❑ District's Title I and non-Title I schools form essentially disjoint strata of students from different populations
- ❑ Estimated teacher effects are lower on average for teachers teaching in Title I schools
 - Roughly 1/2 standard deviation unit lower in Title I Schools
- ❑ Teachers in Title I schools are much more likely to be considered below average (11 to 25% vs. 0 to 8%)
- ❑ Teachers in Title I schools are less likely to be considered above average (17 to 22% vs. 23 to 58%)

Including Student Covariates in the Model Is Not Necessarily Going to Work

- ❑ Available covariates might be incomplete
- ❑ Teacher effects are based on residuals
- ❑ Model attributes all factors correlated with covariates to covariates
 - Effects of covariate
 - Teacher effectiveness correlated with covariates
- ❑ Model over adjusts
- ❑ Estimated effects remain confounded by student characteristics

Partial Fixes for Covariate Modeling Exist

- Using within classroom variance on covariates can adjust for student level covariates without over adjusting
- Does not work for school or classroom level covariates
 - Neighborhood effects
 - Contextual effects

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How Sensitive Are Estimated Effects to Assumptions About Missing Data?

- ❑ **Missing data are pervasive**
- ❑ **In urban data set only 20% of observations are complete**
 - **6,417 first graders in 1998**
 - **370 repeated a grade**
 - **29 skipped a grade**
 - **2,615 transferred out of the district**
 - **317 missed testing**
 - **3,915 students enter the cohort after grade 1**
- ❑ **Observations with missing scores are also missing teacher links**
- ❑ **Must make assumptions about missing links and scores**

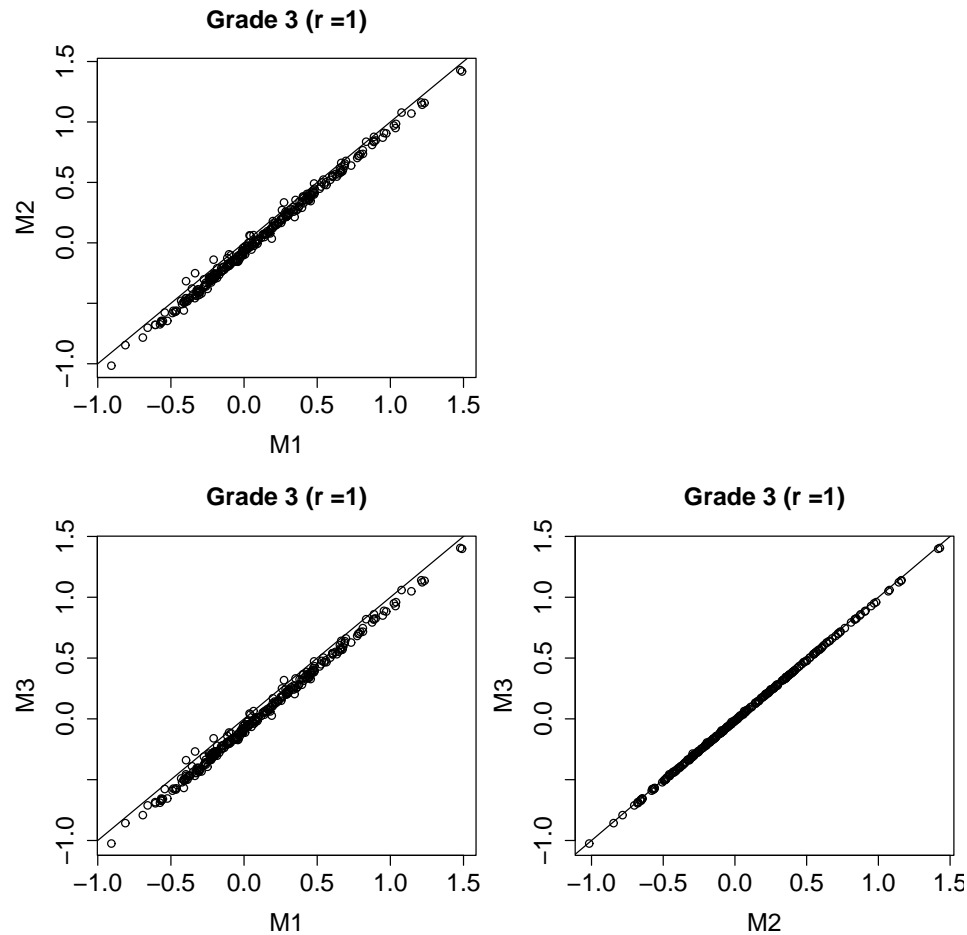
Missing Links Must Be Imputed

- Bayesian method uses data augmentation to find posteriors given the observed data and priors
- Teacher links are required for augmentation
- Values of missing links matter for years prior to the last year of observed data
 - Persistence assumes prior teachers contribute directly to current year scores
 - After the last observed score, values of the missing links do not matter and there is no information about these teachers in the data, we set these links to a teacher with zero effect

Methods to Missing Links

- **Method 1. Set all missing links to a teacher with zero effect**
- **Method 2. Give each student a “pseudo-teacher”**
 - **Pseudo-teacher links only to this one student**
 - **Estimate pseudo-teacher effects**
 - **Variance of pseudo-teachers effects is assumed to equal that of other teachers from the same grade**
- **Method 3. Give each student a “pseudo-teacher” but let pseudo-teachers have different variance components from other teachers**

Teacher Effect Estimates Nearly Invariant to Method for Missing Link



Simulation Study of Sensitivity to Assumption That Scores are Missing At Random (MAR)

- Generate data from persistence model with no covariates and no school effects
- All the α parameters equal one (i.e., the layered model)
- $\epsilon_{it} = \sigma_t(\delta_i + \zeta_{it})$
 - $\delta_i \sim N(0, .7), \nu^2 = .7$
 - $\zeta_{it} \sim N(0, .3)$
 - σ_t^2 based on the marginal variance of our urban test data

Simulation Study Missing Data Mechanism

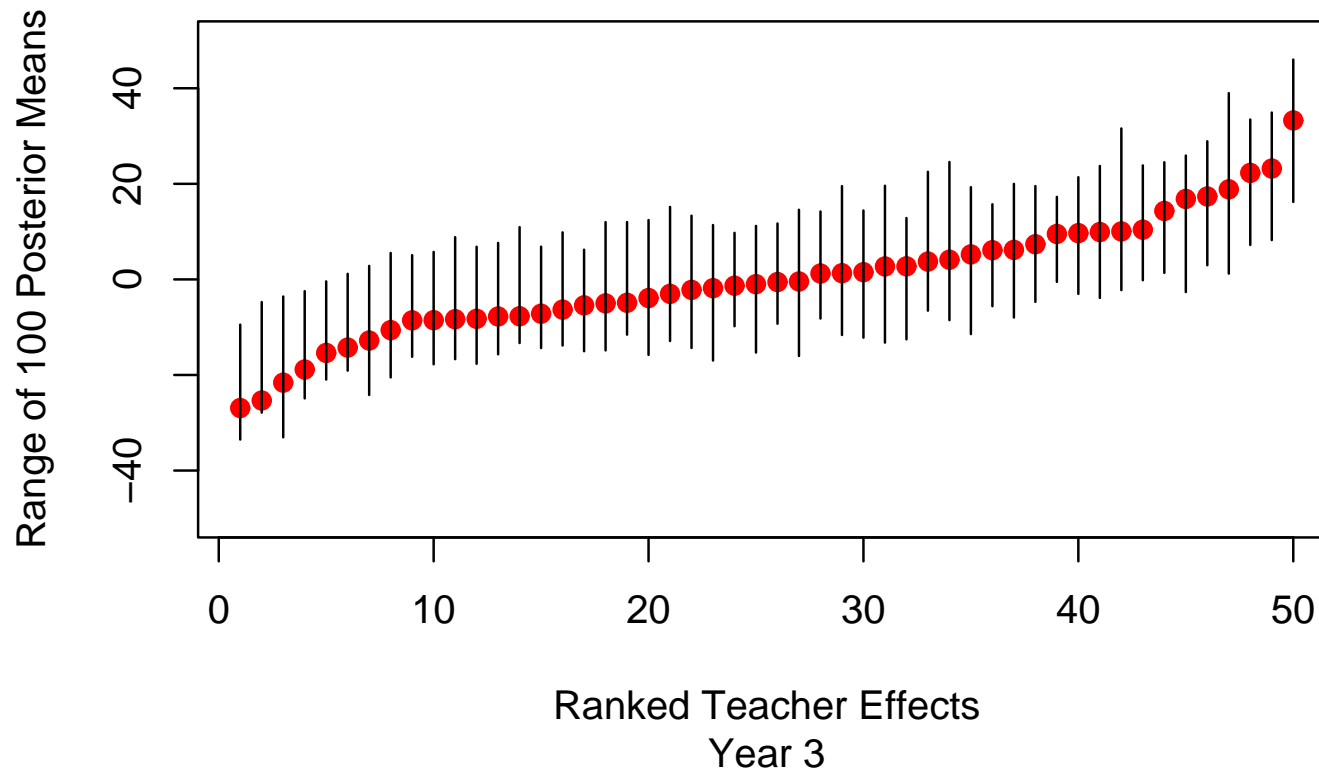
- Probability of missing depends on student's δ value
 - Number of observed year T_{obs} depends on δ through the model

$$P(T_{obs} \geq t) = \frac{1}{1 + e^{\mu t - \beta \delta}}, \text{ for } t = 1, \dots, 5 / \quad (1)$$

- $\beta > 0$ – students with lower values of δ are more likely to have more missing scores
- Pattern of missing is selected at random
- Probability of missing scores matches our urban test score data
- 5 grades, 50 teachers per grade, 25 students per teacher

Teacher Effects Are Relatively Robust to Violation of MAR Assumption

Recovery of True Teacher Effects
Layered Model -- NI Missing Data

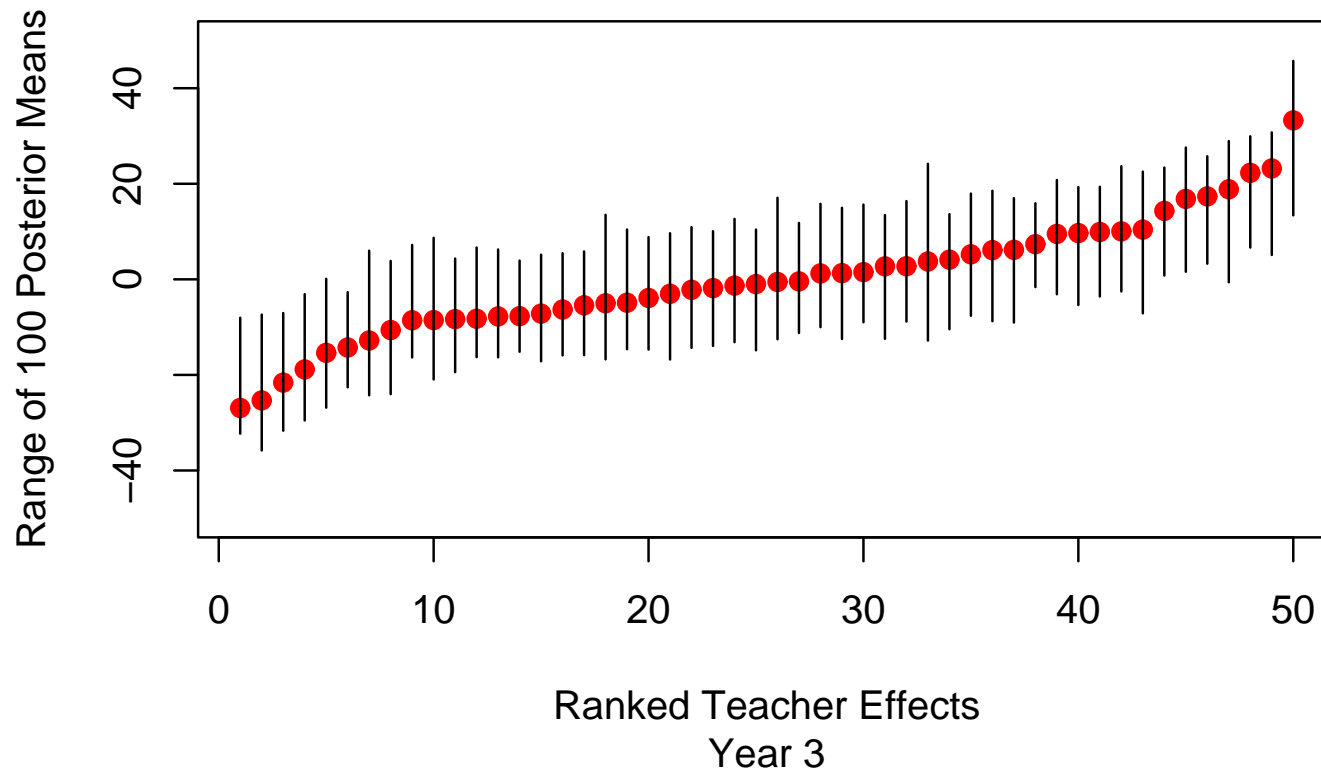


But There is Bias for Teachers at the Low End of the Sample and In the Overall Average Effect

- Estimates tend to be above the true value for teachers at the low end of the sample**
- Average estimated teacher effect estimate is greater than zero for grades 1 to 3**
 - For grades 1 to 5, means are 34.1, 22.5, 9.5, .2, and -.1 % of standard deviation unit**
- Bias can result in misclassifying teachers as above or below zero**
- Bias occurs only when data are not MAR**
- Bias occurs only when we are missing teacher links**

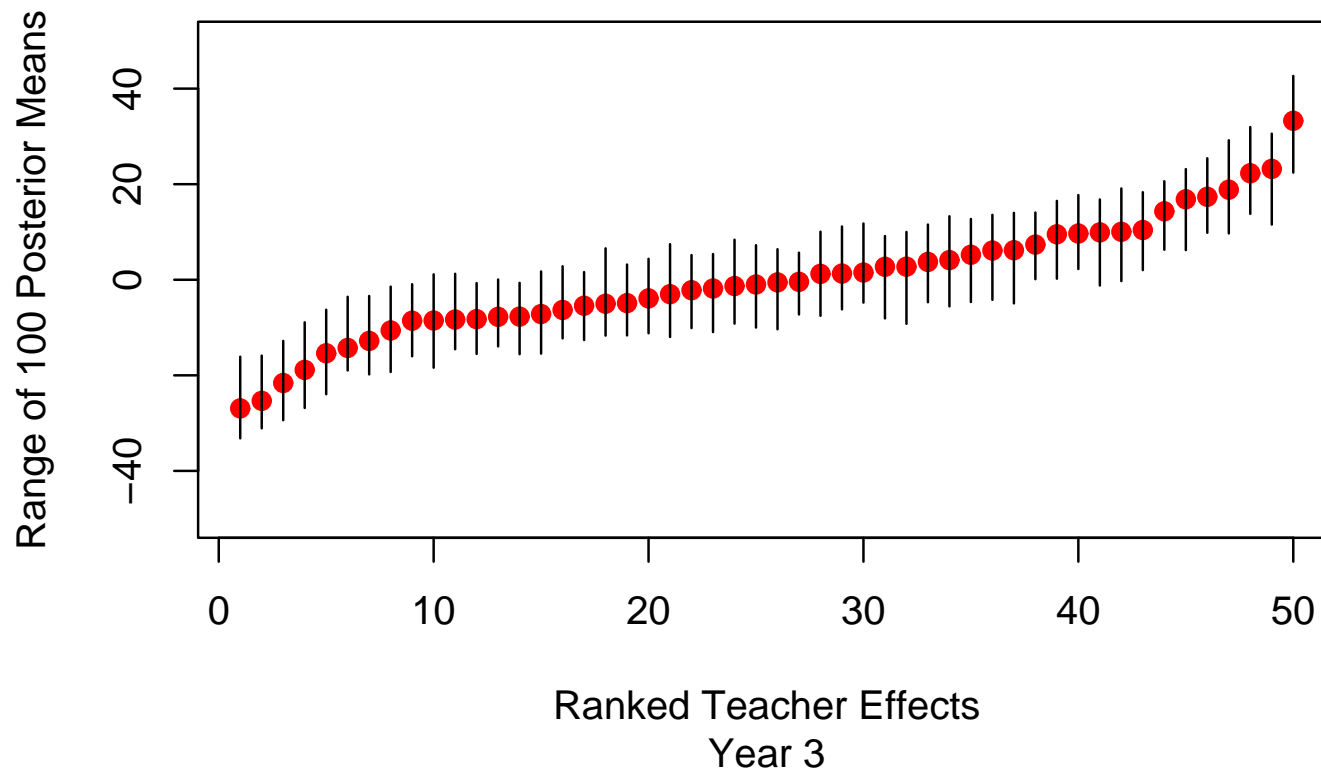
Teacher Effects When Data are MAR

Recovery of True Teacher Effects
Layered Model -- MAR Missing Data



Teacher Effects When Data are Complete

Recovery of True Teacher Effects
Layered Model -- NO Missing Data



Outline

- Persistence of teacher effects
- Effects of student background variables
- Missing scores and missing teacher links
- **Modeling approach**

Are the Complex Models Worth the Effort?

- ❑ Persistence model, layered model, and cross-classified models lack transparency
- ❑ We investigate the estimators and compare average mean square error for several approaches.

Tractable Simple Case for Analytic Investigations

$$y_{i1} = \mu_1 + \phi'_{i1} \theta_1 + \epsilon_{i1}$$

$$y_{i2} = \mu_2 + \alpha_{21} \phi'_{i1} \theta_1 + \phi'_{i2} \theta_2 + \epsilon_{i2}$$

- Complete data
- Known parameters
- No year one clustering

$$y_{i1} = \mu_1 + \epsilon_{i1}$$

$$y_{i2} = \mu_2 + \theta_2 + \epsilon_{i2}$$

- Σ equals the variance covariance matrix for $(\epsilon_{i1}, \epsilon_{i2})$
- $d_i = y_{i2} - y_{i1}$

Seven Approaches to Modeling This Data

1. Fixed effects using y_2
2. Fixed effects using d
3. Random effects modeling y_2 ignoring y_1
4. Random effects modeling d ignoring y_1
5. Random effects modeling y_2 and y_1 jointly
6. Random effects modeling d and y_1 jointly
7. Random effects with y_1 as a covariate in the model for y_2

Teacher Effects Estimated by the BLUP

- The BLUP, $\hat{\theta}$, is given by

$$\hat{\theta} = (\mathbf{Z}'\mathbf{R}^{-1}\mathbf{Z} + \mathbf{D}^{-1})^{-1} \mathbf{Z}'\mathbf{R}^{-1}\mathbf{r}$$

- \mathbf{R} – the block-diagonal covariance matrix for the entire vector of residual error terms
- \mathbf{Z} – the matrix of links from student scores to teachers
- \mathbf{D} – covariance matrix for the random teacher effects ($\tau\mathbf{I}$)
- \mathbf{r} – the vector of estimated residuals $y - \mu$

Evaluating the BLUP

- Evaluating $Z'R^{-1}r$
 - Elements are proportional to classroom means of adjusted residuals, e
 - Residuals determined by the R^{-1} for each approach
- Evaluating $(Z'R^{-1}Z + D^{-1})^{-1}$
 - The matrix is diagonal
 - Elements are proportional to a “shrinkage factor” λ
 - For random effects, $\lambda = \tau / (\tau + \nu^2)$
 - ν^2 is the variance of the adjusted residual mean conditional on θ
 - For fixed effects, $\lambda = 1$
- $\hat{\theta} = \lambda \bar{e}$
- $MSE = \lambda \nu^2$

Joint Modeling Provides Smallest MSE of Seven Alternatives

Approach	MSE ($\lambda \times \nu^2$)
Fixed Effects with y_2	$1 \times \sigma_2^2/n$
Fixed Effects with d	$1 \times \sigma_d^2/n$
Random Effects, y_2 Alone	$\left(\frac{\tau}{\tau + \sigma_2^2/n} \right) \times \sigma_2^2/n$
Random Effects, d Alone	$\left(\frac{\tau}{\tau + \sigma_d^2/n} \right) \times \sigma_d^2/n$
Joint Modeling (y_1, y_2) or (y_1, d) y_1 as covariate	$\left(\frac{\tau}{\tau + \sigma_2^2(1 - \rho^2)/n} \right) \times \sigma_2^2(1 - \rho^2)/n$

With Clustering in Both Years Joint Modeling Adjusts for Year 1 Teachers

- **Step 1. Generate preliminary year 1 teacher effect estimates ignoring year 2 teacher assignments**
 - **Similar to Approach 3, estimate residuals from regressing estimates of ϵ_1 on ϵ_2 , create classroom means of the residuals, and shrink toward zero**
- **Step 2. Adjust $r_1 = y_1 - \mu_1$ by the estimated year 1 teacher effects to obtain \tilde{r}_1**
- **Step 3. Adjust $r_2 = y_2 - \mu_2$ by $\beta_{21}\tilde{r}_1$**
- **Step 4. Average the adjusted year two residuals by classrooms and shrink**

Joint Modeling Is Preferable When Clustering Exits in Year 1

- When class assignments are completely balanced, teacher effect estimators equal those that ignore clustering in year 1
- When not balanced, the estimator is not equivalent to using y_1 as covariate and should be more efficient because it adjusts for year 1 teachers and it uses correct adjustment for year 1 residuals

What Have We Learned About the Persistence of Teacher Effects?

- ❑ **Teacher effects dampen over time in a large school district**
- ❑ **Inferences can be sensitive to model assumptions about dampening**
 - **Persistence model finds more teachers distinct from zero**
- ❑ **Bayesian implementation improves computation and makes the model feasible even for large problems**

What Have We Learned About the Effects of Student Background Variables?

- ❑ **Linear covariate adjustments can create bias by over adjusting**
- ❑ **We cannot assume joint modeling of multiple scores will remove effects of student background variables**
 - **Particularly concerned when combining segregated populations where gains might differ across groups**
- ❑ **Best approach might be to compare teachers teaching similar populations**

What Have We Learned About Missing Data?

- Missing teacher links are a challenge to modeling incomplete data
- Ordering of teachers appears robust to approach to missing links and to violations of MAR
- Inferences about teachers being above or below average might be sensitive to violations of MAR
 - Centering estimated effects might provide an ad hoc fix to bias
- Further studies should:
 - Explore alternative teacher assignment mechanisms
 - Explore missing data mechanisms that depend on gains
 - Explore the effect of missing data when α s are less than one

What Have We Learned About Complex Models?

- **Joint modeling is doing regression adjustments**
 - **In simple special cases joint modeling is identical to covariate adjustment with year one score**
- **Joint modeling of scores is likely to be more efficient than alternatives**
 - **Uses prior year score efficiently and adjusts for other year teachers**
- **More work needs to be done to explore gains in efficiency from joint modeling**
- **More work needs to be done to see if gains for joint modeling hold when the data are very incomplete and assumptions are violated**